| Notations and Useful Data |  |
| :---: | :---: |
| $\mathbb{N}$ | The set of positive integers |
| $\mathbb{R}$ | The set of real numbers |
| $\mathbb{R}^{n}$ | $\left\{\left(x_{1}, x_{2}, \ldots, x_{n}\right): x_{i} \in \mathbb{R}, i=1,2, \ldots, n\right\}, n=2,3, \ldots$ |
| $\ln x$ | Natural logarithm of $x, x>0$ |
| $\operatorname{det}(M)$ | Determinant of a square matrix $M$ |
| adj M | Adjoint of a square matrix $M$, that is, transpose of cofactor matrix of $M$ |
| $\emptyset$ | Empty set |
| $E^{C}$ | Complement of event $E$ |
| $P(E)$ | Probability of event $E$ |
| $P(E \mid F)$ | Conditional probability of event $E$ given the occurrence of event $F$ |
| $E(X)$ | Expectation of a random variable $X$ |
| $\operatorname{Var}(X)$ | Variance of a random variable $X$ |
| $\operatorname{Cov}(X, Y)$ | Covariance between random variables $X$ and $Y$ |
| $\operatorname{Bin}(n, p)$ | Binomial distribution with parameters $n$ and $p, n \in \mathbb{N}, 0<p<1$ |
| $U(a, b)$ | Continuous uniform distribution on the interval ( $a, b$ ), a<b,a,b, 顺 |
| $\operatorname{Exp}(\lambda)$ | Exponential distribution with the probability density function $f(x)=\left\{\begin{array}{cl} \lambda e^{-\lambda x}, & \text { if } x>0 \\ 0, & \text { if } x \leq 0 \end{array}, \text { for } \lambda>0\right.$ |
| $N\left(\mu, \sigma^{2}\right)$ | Normal distribution with mean $\mu$ and variance $\sigma^{2}, \mu \in \mathbb{R}, \sigma>0$ |
| $N_{2}\left(\mu_{1}, \mu_{2}, \sigma_{1}^{2}, \sigma_{2}^{2}, \rho\right)$ | Bivariate normal distribution with means $\mu_{1}, \mu_{2}$, variances $\sigma_{1}^{2}, \sigma_{2}^{2}$ and correlation $\rho, \quad \mu_{1} \in \mathbb{R}, \mu_{2} \in \mathbb{R}, \sigma_{1}>0, \sigma_{2}>0,-1<\rho<1$ |
| $\phi(\cdot)$ | The probability density function of $N(0,1)$ random variable |
| $\Phi(\cdot)$ | The cumulative distribution function of $N(0,1)$ random variable |
| $\chi_{n}^{2}$ | Central chi-square distribution with $n$ degrees of freedom, $n=1,2, \ldots$ |
| $t_{n}$ | Central Student's $t$ distribution with $n$ degrees of freedom, $n=1,2, \ldots$ |
| $F_{m, n}$ | Snedecor's central $F$-distribution with $m$ and $n$ degrees of freedom, $m, n \in \mathbb{N}$ |
| $\chi_{n, \alpha}^{2}$ | A constant such that $P\left(X>\chi_{n, \alpha}^{2}\right)=\alpha$, where $X$ has central chi-square distribution with $n$ degrees of freedom, $n=1,2, \ldots ; \alpha \in(0,1)$ |
| $\overline{t_{n, \alpha}}$ | A constant such that $P\left(X>t_{n, \alpha}\right)=\alpha$, where $X$ has central Student's $t$ distribution with $n$ degrees of freedom, $n=1,2, \ldots ; \alpha \in(0,1)$ |
| $\xrightarrow{\text { d }}$ | Convergence in distribution |
| $\xrightarrow{P}$ | Convergence in probability |
| i. i. d. | Independent and identically distributed |
|  |  |
|  |  |

## Section A: Q. 1 - Q. 10 Carry ONE mark each.

| Q.1 | Let $a_{n}=1+\frac{1}{2}+\cdots+\frac{1}{n}$ and $b_{n}=\frac{n^{2}}{2^{n}}$ for all $n \in \mathbb{N}$. Then |
| ---: | :--- |
| (A) | $\left\{a_{n}\right\}$ is a Cauchy sequence but $\left\{b_{n}\right\}$ is NOT a Cauchy sequence |
| (B) | $\left\{a_{n}\right\}$ is NOT a Cauchy sequence but $\left\{b_{n}\right\}$ is a Cauchy sequence |
| (C) | both $\left\{a_{n}\right\}$ and $\left\{b_{n}\right\}$ are Cauchy sequences |
| (D) | neither $\left\{a_{n}\right\}$ nor $\left\{b_{n}\right\}$ is a Cauchy sequence |
| Q.2 | Let $f(x, y)=2 x^{4}-3 y^{2}$ for all $(x, y) \in \mathbb{R}^{2}$. Then |
| (A) | $f$ has a point of local minimum |
| (B) | $f$ has a point of local maximum |
| (C) | $f$ has a saddle point no point of local minimum, no point of local maximum, and no saddle point |
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| Q.3 | Let $A=\left(\begin{array}{ll}a & 0 \\ c & d\end{array}\right)$ be a real matrix, where $a d=1$ and $c \neq 0$. If |
| ---: | :--- |
|  | $A^{-1}+(\operatorname{adj} A)^{-1}=\left(\begin{array}{ll}\alpha & \beta \\ \gamma & \delta\end{array}\right)$, |
| (hen $(\alpha, \beta, \gamma, \delta)$ is equal to |  |
| (B) | $(a+d, 0,0, a+d)$ |
| (C) | $(a, 0,0, d)$ |
| (D) | $(a, 0, c, d)$ |
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| Q.4 | A bag has 5 blue balls and 15 red balls. Three balls are drawn at random from the <br> bag simultaneously. Then the probability that none of the chosen balls is blue <br> equals |
| ---: | :--- |
| (A) | $\frac{75}{152}$ |
| (B) | $\frac{91}{228}$ |
| (C) | $\frac{27}{64}$ |
| (D) | $\frac{273}{800}$ |
| (D) | $\xi_{0.25}=2$ |
| (C) | $\xi_{0.25} \geq 4$ |
| (B) | $\xi_{0.75} \leq 4$ |
| (A) | $\xi_{0.75} \geq 5$ <br> Let $Y$ be a continuous random variable such that $P(Y>0)=1$ and $E(Y)=1$. <br> random variable $Y . ~ T h e n ~ w h i c h ~ o f ~ t h e ~ f o l l o w i n g ~ s t a t e m e n t s ~ i s ~ a l w a y s ~ c o r r e c t ? ~$ |
| let $\xi_{p}$ denote the $p^{\text {th }}$ quantile of the probability distribution of the |  |


| Q.6 | Let $X$ be a continuous random variable having the $U(-2,3)$ distribution. Then <br> which of the following statements is correct? |
| ---: | :--- |
| (A) | $2 X+5$ has the $U(1,10)$ distribution |
| (B) | $7-6 X$ has the $U(-11,19)$ distribution |
| (C) | $3 X^{2}+5$ has the $U(5,32)$ distribution |
| (D) | $1 X \mid$ has the $U(0,3)$ distribution |
| (D) | $\frac{3}{2} e^{-1}$ |
| (C) | $\frac{5}{2} e^{-1}$ |
| (B) | Let $X$ be a random variable having the Poisson distribution with mean 1. Let <br> $g: \mathbb{N} \cup\{0\} \rightarrow \mathbb{R}$ be defined by <br> (A) |
| -1 |  |



| Q.10 | Let $x_{1}, x_{2}, x_{3}, x_{4}$ be the observed values of a random sample from a $N\left(\mu, \sigma^{2}\right)$ <br> distribution, where $\mu \in \mathbb{R}$ and $\sigma \in(0, \infty)$ are unknown parameters. Let $\bar{x}$ and <br> $s=\sqrt{\frac{1}{3} \sum_{i=1}^{4}\left(x_{i}-\bar{x}\right)^{2}}$ be the observed sample mean and the sample standard <br> deviation, respectively. For testing $H_{0}: \mu=0$ against $H_{1}: \mu \neq 0$, the likelihood <br> ratio test of size $\alpha=0.05$ rejects $H_{0}$ if and only if $\frac{\|\bar{x}\|}{s}>k$. Then the value of $k$ <br> is |
| ---: | :--- |
| (A) | $\frac{1}{2} t_{3,0.025}$ |$|$| (B) | $t_{3,0.025}$ |
| :--- | :--- |
| (C) | $2 t_{3,0.05}$ |
| (D) | $\frac{1}{2} t_{3,0.05}$ |
|  |  |
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$\square$
Section A: Q. 11 - Q. 30 Carry TWO marks each.

| Q.11 | For $n \in \mathbb{N}$, let $a_{n}=\sqrt{n} \sin ^{2}\left(\frac{1}{n}\right) \cos n$, and $b_{n}=\sqrt{n} \sin \left(\frac{1}{n^{2}}\right) \cos n$. Then |
| ---: | :--- |
| (A) | the series $\sum_{n=1}^{\infty} a_{n}$ converges but the series $\sum_{n=1}^{\infty} b_{n}$ does NOT converge |
| (B) | the series $\sum_{n=1}^{\infty} a_{n}$ does NOT converge but the series $\sum_{n=1}^{\infty} b_{n}$ converges |
| (C) | both the series $\sum_{n=1}^{\infty} a_{n}$ and $\sum_{n=1}^{\infty} b_{n}$ converge |
| (D) | neither the series $\sum_{n=1}^{\infty} a_{n}$ nor the series $\sum_{n=1}^{\infty} b_{n}$ converges |
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| Q.13 | Let $f(x, y)=\|x y\|+x$ for all $(x, y) \in \mathbb{R}^{2}$. Then the partial derivative of $f$ with <br> respect to $x$ exists |
| ---: | :--- |
| (A) | at $(0,0)$ but NOT at $(0,1)$ |
| (B) | at $(0,1)$ but NOT at $(0,0)$ |
| (C) | at $(0,0)$ and $(0,1)$, both |
| (D) | neither at $(0,0)$ nor at $(0,1)$ |
| Q.14 | Let $f(x)=4 x^{2}-\sin x+\cos 2 x$ for all $x \in \mathbb{R}$. Then $f$ has |
| (A) | a point of local maximum |
| (B) | no point of local minimum |
| (C) | at least two points of local minima one point of local minimum |
|  |  |


| Q.15 | Consider the improper integrals |
| ---: | :--- |
|  | $I_{1}=\int_{1}^{\infty} \frac{t \sin t}{e^{t}} d t$ and $I_{2}=\int_{1}^{\infty} \frac{1}{\sqrt{t}} \ln \left(1+\frac{1}{t}\right) d t$ |
| (A) | $I_{1}$ converges but $I_{2}$ does NOT converge |
| (B) | $I_{1}$ does NOT converge but $I_{2}$ converges |
| (C) | both $I_{1}$ and $I_{2}$ converge |
| (D) | neither $I_{1}$ nor $I_{2}$ converges |
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| Q. 17 | Let $\Omega=\{1,2,3,4,5,6\}$. Then which of the following classes of sets is an algebra? |
| :---: | :---: |
| (A) | $\mathcal{F}_{1}=\{\emptyset, \Omega,\{1,2\},\{3,4\},\{3,6\}\}$ |
| (B) | $\mathcal{F}_{2}=\{\varnothing, \Omega,\{1,2,3\},\{4,5,6\}\}$ |
| (C) | $\mathcal{F}_{3}=\{\emptyset, \Omega,\{1,2\},\{4,5\},\{1,2,4,5\},\{3,4,5,6\},\{1,2,3,6\}\}$ |
| (D) | $\mathcal{F}_{4}=\{\emptyset,\{4,5\},\{1,2,3,6\}\}$ |
| Q. 18 | Two fair coins $S_{1}$ and $S_{2}$ are tossed independently once. Let the events $E, F$ and $G$ be defined as follows: <br> $E$ : Head appears on $S_{1}$ <br> $F$ : Head appears on $S_{2}$ <br> $G$ : The same outcome (head or tail) appears on both $S_{1}$ and $S_{2}$ <br> Then which of the following statements is NOT correct? |
| (A) | $E$ and $F$ are independent |
| (B) | $F$ and $G$ are independent |
| (C) | $E$ and $G^{C}$ are independent |
| (D) | $E, F$, and $G$ are mutually independent |


| Q.19 | Let $f_{1}(x)$ be the probability density function of the $N(0,1)$ distribution and $f_{2}(x)$ <br> be the probability density function of the $N(0,6)$ distribution. Let $Y$ be a random <br> variable with probability density function <br> Then $\operatorname{Var}(Y)$ is equal to |
| ---: | :--- |
| (A) | 7 |
| (B) | 3 |
| (C) | 3.5 |
| (D) | 1 |
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| Q. 20 | Which of the following functions represents a cumulative distribution function? |
| :---: | :---: |
| (A) | $F_{1}(x)=\left\{\begin{array}{cc} 0, & \text { if } x<\frac{\pi}{4} \\ \sin x, & \text { if } \frac{\pi}{4} \leq x<\frac{3 \pi}{4} \\ 1, & \text { if } x \geq \frac{3 \pi}{4} \end{array}\right.$ |
| (B) | $F_{2}(x)=\left\{\begin{array}{cc} 0, & \text { if } x<0 \\ 2 \sin x, & \text { if } 0 \leq x<\frac{\pi}{4} \\ 1, & \text { if } x \geq \frac{\pi}{4} \end{array}\right.$ |
| (C) | $F_{3}(x)=\left\{\begin{array}{cc} 0, & \text { if } x<0 \\ x, & \text { if } 0 \leq x<\frac{1}{3} \\ x+\frac{1}{3}, & \text { if } \frac{1}{3} \leq x \leq \frac{1}{2} \\ 1, & \text { if } x>\frac{1}{2} \end{array}\right.$ |
| (D) | $F_{4}(x)=\left\{\begin{array}{cc} 0, & \text { if } x<0 \\ \sqrt{2} \sin x, & \text { if } 0 \leq x<\frac{\pi}{4} \\ 1, & \text { if } x \geq \frac{\pi}{4} \end{array}\right.$ |
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| Q.21 | Let $X$ be a random variable such that $X$ and $-X$ have the same distribution. <br> Let $Y=X^{2}$ be a continuous random variable with the probability density function |
| ---: | :--- |
| $\qquad$Then $E\left((X-1)^{4}\right)$ is equal to$\frac{e^{-\frac{y}{2}}}{\sqrt{2 \pi y}}$, if $y>0$. <br> 0, if $y \leq 0$  <br> (A) 9 <br> (C) 10 <br> (D) 12 <br>   <br>   |  |


| Q.22 | Suppose that random variable $X$ has $\operatorname{Exp}\left(\frac{1}{5}\right)$ distribution and, for any $x>0$, the <br> conditional distribution of random variable $Y$, given $X=x$, is $N(x, 2)$. Then <br> $\operatorname{Var}(X+Y)$ is equal to |
| ---: | :--- |
| (A) | 52 |
| (B) | 50 |
| (C) | 2 |
| (D) | 102 |
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| Q.23 | Let the random vector $(X, Y)$ have the joint probability density function |
| ---: | :--- |
| $\qquad f(x, y)= \begin{cases}\frac{1}{x}, & \text { if } 0<y<x<1 . \\ 0, & \text { otherwise }\end{cases}$ |  |
| (A) | $\frac{1}{6}$ |
| (B) | $\frac{1}{12}$ |
| (C) | $\frac{1}{18}$ |
| (D) | $\frac{1}{24}$ |
|  |  |


| Q. 24 | Let $\left(X_{1}, Y_{1}\right),\left(X_{2}, Y_{2}\right), \ldots,\left(X_{20}, Y_{20}\right)$ be a random sample from the $N_{2}\left(0,0,1,1, \frac{3}{4}\right)$ distribution. Define $\bar{X}=\frac{1}{20} \sum_{i=1}^{20} X_{i}$ and $\bar{Y}=\frac{1}{20} \sum_{i=1}^{20} Y_{i}$. Then $\operatorname{Var}(\bar{X}-\bar{Y})$ is equal to |
| :---: | :---: |
| (A) | $\frac{1}{16}$ |
| (B) | $\frac{1}{40}$ |
| (C) | $\frac{1}{10}$ |
| (D) | $\frac{3}{40}$ |
|  |  |
|  |  |
|  |  |
|  | $1$ |
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| Q.25 | For $n \in \mathbb{N}$, let $X_{n}$ be a random variable having the $\operatorname{Bin}\left(n, \frac{1}{4}\right)$ distribution. Then |
| ---: | :--- |
|  | $\lim _{n \rightarrow \infty}\left[P\left(X_{n} \leq \frac{2 n-\sqrt{3 n}}{8}\right)+P\left(\frac{n}{6} \leq X_{n} \leq \frac{n}{3}\right)\right]$ <br> $($ You may use $\Phi(0.5)=0.6915, ~$ <br> $0.9772)$ |
| (A) | 1.6915 |
| (B) | 1.3085 |
| (C) | 1.1587 |
| (D) | 0.6915 |
|  |  |


| Q.26 | Let $X_{1}, X_{2}, \ldots, X_{10}$ be a random sample from the $N(3,4)$ distribution and let <br> $Y_{1}, Y_{2}, \ldots, Y_{15}$ be a random sample from the $N(-3,6)$ distribution. Assume that <br> the two samples are drawn independently. Define |
| ---: | :--- |
|  | $\bar{X}=\frac{1}{10} \sum_{i=1}^{10} X_{i}, \bar{Y}=\frac{1}{15} \sum_{j=1}^{15} Y_{j}$, and $S=\sqrt{\frac{1}{9} \sum_{i=1}^{10}\left(X_{i}-\bar{X}\right)^{2}}$. |
| (A) | $N\left(0, \frac{4}{5}\right)$ |
| (B) | $\chi_{9}^{2}$ |
| (C) | $t_{9}$ |
| (D) | $t_{23}$ |
|  |  |
|  |  |


| Q. 27 | For $n \geq 2$, let $\epsilon_{1}, \epsilon_{2}, \ldots, \epsilon_{n}$ be i.i.d. random variables having the $N(0,1)$ distribution. Consider $n$ independent random variables $Y_{1}, Y_{2}, \ldots, Y_{n}$ defined by $Y_{i}=\beta i+\epsilon_{i}, \quad i=1,2, \ldots, n$ <br> where $\beta \in \mathbb{R}$. Define $\bar{Y}=\frac{1}{n} \sum_{i=1}^{n} Y_{i}, T_{1}=\frac{2 \bar{Y}}{n+1}$, and $T_{2}=\frac{1}{n} \sum_{i=1}^{n} \frac{Y_{i}}{i}$. Then which of the following statements is NOT correct? |
| :---: | :---: |
| (A) | $T_{1}$ is an unbiased estimator of $\beta$ |
| (B) | $T_{2}$ is an unbiased estimator of $\beta$ |
| (C) | $\operatorname{Var}\left(T_{1}\right)<\operatorname{Var}\left(T_{2}\right)$ |
| (D) | $\operatorname{Var}\left(T_{1}\right)=\operatorname{Var}\left(T_{2}\right)$ |
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| Q. 28 | A biased coin, with probability of head as $p$, is tossed $m$ times independently. It is known that $p \in\left\{\frac{1}{4}, \frac{3}{4}\right\}$ and $m \in\{3,5\}$. If 3 heads are observed in these $m$ tosses, then which of the following statements is correct? |
| :---: | :---: |
| (A) | $\left(3, \frac{3}{4}\right)$ is a maximum likelihood estimator of ( $m, p$ ) |
| (B) | $\left(5, \frac{1}{4}\right)$ is a maximum likelihood estimator of ( $m, p$ ) |
| (C) | $\left(5, \frac{3}{4}\right)$ is a maximum likelihood estimator of $(m, p)$ |
| (D) | Maximum likelihood estimator of ( $m, p$ ) is NOT unique |
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| Q.29 | Let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample from an $\operatorname{Exp}(\lambda)$ distribution, where <br> $\lambda \in\{1,2\}$. For testing $H_{0}: \lambda=1$ against $H_{1}: \lambda=2$, the most powerful test of <br> size $\alpha, \alpha \in(0,1)$, will reject $H_{0}$ if and only if |
| ---: | :--- |
| (A) | $\sum_{i=1}^{n} X_{i} \leq \frac{1}{2} \chi_{2 n, 1-\alpha}^{2}$ |
| (B) | $\sum_{i=1}^{n} X_{i} \geq 2 \chi_{2 n, 1-\alpha}^{2}$ |
| (C) | $\sum_{i=1}^{n} X_{i} \leq \frac{1}{2} \chi_{n, 1-\alpha}^{2}$ |
| (D) | $\sum_{i=1}^{n} X_{i} \geq 2 \chi_{n, 1-\alpha}^{2}$ |
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| Q.30 | Let $X_{1}, X_{2}, \ldots, X_{10}$ be a random sample from a $N\left(0, \sigma^{2}\right)$ distribution, where <br> $\sigma>0$ is unknown. For testing $H_{0}: \sigma^{2} \leq 1$ against $H_{1}: \sigma^{2}>1$, a test of size <br> $\alpha=0.05$ rejects $H_{0}$ if and only if $\sum_{i=1}^{10} X_{i}^{2}>18.307$. Let $\beta$ be the power of this <br> test, at $\sigma^{2}=2$. Then $\beta$ lies in the interval <br> (You may use $\chi_{10,0.05}^{2}=18.307, \chi_{10,0.1}^{2}=15.9872, \chi_{10,0.25}^{2}=12.5489$, <br> $\chi_{10,0.5}^{2}=9.3418, \chi_{10,0.75}^{2}=6.7372, \chi_{10,0.9}^{2}=4.8652, \chi_{10,0.95}^{2}=3.9403$, <br> $\left.\chi_{10,0.975}^{2}=3.247\right)$ |
| ---: | :--- |
| (A) | $(0.50,0.75)$ |
| (B) | $(0.75,0.90)$ |
| (C) | $(0.90,0.95)$ |
| (D) | $(0.95,0.975)$ |
|  |  |
|  |  |



Section B: Q. 31 - Q. 40 Carry TWO marks each.

| Q.31 | Let $a_{1}=1, a_{n+1}=a_{n}\left(\frac{\sqrt{n}+\sin n}{n}\right)$ and $b_{n}=a_{n}^{2}$ for all $n \in \mathbb{N}$. Then which of the <br> following statements is/are correct? |
| ---: | :--- |
| (A) | the series $\sum_{n=1}^{\infty} a_{n}$ converges |
| (B) | the series $\sum_{n=1}^{\infty} b_{n}$ converges |
| (C) | the series $\sum_{n=1}^{\infty} a_{n}$ converges but the series $\sum_{n=1}^{\infty} b_{n}$ does NOT converge |
| (D) | neither the series $\sum_{n=1}^{\infty} a_{n}$ nor the series $\sum_{n=1}^{\infty} b_{n}$ converges |
|  |  |
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| Q. 32 | Let $f: \mathbb{R} \rightarrow \mathbb{R}$ be a twice differentiable function such that $f(0)=0, f(2)=4, f(4)=4 \text { and } f(8)=12 .$ <br> Then which of the following statements is/are correct? |
| :---: | :---: |
| (A) | $f^{\prime}(x) \leq 1$ for all $x \in[0,2]$ |
| (B) | $f^{\prime}\left(x_{1}\right)>1$ for some $x_{1} \in[0,2]$ |
| (C) | $f^{\prime}\left(x_{2}\right)>1$ for some $x_{2} \in[4,8]$ |
| (D) | $f^{\prime \prime}\left(x_{3}\right)=0$ for some $x_{3} \in[0,8]$ |
|  |  |
|  |  |
| Q. 33 | Let $A$ be a $3 \times 3$ real matrix. Suppose that 1 and 2 are characteristic roots of $A$, and 12 is a characteristic root of $A+A^{2}$. Then which of the following statements is/are correct? |
| (A) | $\operatorname{det}(A) \neq 0$ |
| (B) | $\operatorname{det}\left(A+A^{2}\right) \neq 0$ |
| (C) | $\operatorname{det}(A)=0$ |
| (D) | trace of $\left(A+A^{2}\right)$ is 20 |



| Q. 35 | Let $X$ be a continuous random variable with a probability density function $f$ and the moment generating function $M(t)$. Suppose that $f(x)=f(-x)$ for all $x \in \mathbb{R}$ and the moment generating function $M(t)$ exists for $t \in(-1,1)$. Then which of the following statements is/are correct? |
| :---: | :---: |
| (A) | $P(X=-X)=1$ |
| (B) | 0 is the median of $X$ |
| (C) | $M(t)=M(-t)$ for all $t \in(-1,1)$ |
| (D) | $E(X)=1$ |
|  |  |
|  |  |
| Q. 36 | Let $X$ and $Y$ be independent random variables having $\operatorname{Bin}(18,0.5)$ and $\operatorname{Bin}(20,0.5)$ distributions, respectively. Further, let $U=\min \{X, Y\}$ and $V=\max \{X, Y\}$. Then which of the following statements is/are correct? |
| (A) | $E(U+V)=19$ |
| (B) | $E(\|X-Y\|)=E(V-U)$ |
| (C) | $\operatorname{Var}(U+V)=16$ |
| (D) | $38-(X+Y)$ has Bin $(38,0.5)$ distribution |

$\left.\begin{array}{|r|l|}\hline \text { Q.37 } & \begin{array}{l}\text { Let } X \text { and } Y \text { be continuous random variables having the joint probability density } \\ \text { function }\end{array} \\ \qquad f(x, y)=\left\{\begin{array}{cc}e^{-x}, \quad \text { if } 0 \leq y<x<\infty \\ 0, & \text { otherwise }\end{array}\right. \\ \text { Then which of the following statements is/are correct? }\end{array}\right\}$

| Q. 38 | For $n \geq 2$, let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample from a distribution with $E\left(X_{1}\right)=0, \operatorname{Var}\left(X_{1}\right)=1$ and $E\left(X_{1}^{4}\right)<\infty$. Let $\bar{X}_{n}=\frac{1}{n} \sum_{i=1}^{n} X_{i} \text { and } S_{n}^{2}=\frac{1}{n-1} \sum_{i=1}^{n}\left(X_{i}-\bar{X}_{n}\right)^{2}$ <br> Then which of the following statements is/are always correct? |
| :---: | :---: |
| (A) | $E\left(S_{n}^{2}\right)=1$ for all $n \geq 2$ |
| (B) | $\sqrt{n} \bar{X}_{n} \xrightarrow{d} Z$ as $n \rightarrow \infty$, where $Z$ has the $N(0,1)$ distribution |
| (C) | $\bar{X}_{n}$ and $S_{n}^{2}$ are independently distributed for all $n \geq 2$ |
| (D) | $\frac{1}{n} \sum_{i=1}^{n} X_{i}^{2} \xrightarrow{P} 2, \text { as } n \rightarrow \infty$ |
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| Q. 39 | Let $X_{1}, X_{2}, \ldots, X_{50}$ be a random sample from a $N\left(0, \sigma^{2}\right)$ distribution, where $\sigma>0$. Define $\begin{gathered} \bar{X}_{e}=\frac{1}{25} \sum_{i=1}^{25} X_{2 i}, \quad \bar{X}_{o}=\frac{1}{25} \sum_{i=1}^{25} X_{2 i-1}, \\ S_{e}=\sqrt{\frac{1}{24} \sum_{i=1}^{25}\left(X_{2 i}-\bar{X}_{e}\right)^{2}} \text { and } S_{o}=\sqrt{\frac{1}{24} \sum_{i=1}^{25}\left(X_{2 i-1}-\bar{X}_{o}\right)^{2} .} \end{gathered}$ <br> Then which of the following statements is/are correct? |
| :---: | :---: |
| (A) | $\frac{5 \bar{X}_{e}}{s_{e}}$ has $t_{24}$ distribution |
| (B) | $\frac{5\left(\bar{X}_{e}+\bar{X}_{o}\right)}{\sqrt{S_{e}^{2}+S_{o}^{2}}}$ has $t_{49}$ distribution |
| (C) | $\frac{49 S_{o}^{2}}{\sigma^{2}}$ has $\chi_{49}^{2}$ distribution |
| (D) | $\frac{s_{o}^{2}}{s_{e}^{2}}$ has $F_{24,24}$ distribution |
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| Q.40 | Let $\theta_{0}$ and $\theta_{1}$ be real constants such that $\theta_{1}>\theta_{0}$. Suppose that a random sample <br> is taken from a $N(\theta, 1)$ distribution, $\theta \in \mathbb{R}$. For testing $H_{0}: \theta=\theta_{0}$ against <br> $H_{1}: \theta=\theta_{1}$ at level 0.05, let $\alpha$ and $\beta$ denote the size and the power, respectively, <br> of the most powerful test, $\psi_{0}$. Then which of the following statements is/are <br> correct? |
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| (A) | $\beta<\alpha$ |
| (B) | The test $\psi_{0}$ is the uniformly most powerful test of level $\alpha$ for testing $H_{0}: \theta=\theta_{0}$ <br> against $H_{1}: \theta>\theta_{0}$ |
| (C) | $\alpha<\beta$ |
| The test $\psi_{0}$ is the uniformly most powerful test of level $\alpha$ for testing $H_{0}: \theta=\theta_{0}$ |  |
| against $H_{1}: \theta<\theta_{0}$ |  |
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## Section C: Q. 41 - Q. 50 Carry ONE mark each.

| Q.41 | The radius of convergence of the power series $\sum_{n=0}^{\infty} \frac{2 n(x+3)^{n}}{5^{n}}$ is equal to <br> (answer in integer) |
| :--- | :--- |
| Q.42 | Let $f(x)=\int_{-1}^{x^{2}-2 x} e^{t^{2}-t} d t$ for all $x \in \mathbb{R}$. If $f$ is decreasing on (0, $m$ ) and <br> increasing on $(m, \infty)$, then the value of $m$ is equal to <br> integer) |
|  | Qanswer in |
|  | Let $V=\left\{\left(x_{1}, x_{2}, x_{3}, x_{4}\right) \in \mathbb{R}^{4}: x_{1}=x_{2}\right\}$. Consider $V$ as a subspace of $\mathbb{R}^{4}$ over <br> the real field. Then the dimension of $V$ is equal to |
|  | (answer in integer) |


| Q. 44 | If 12 fair dice are independently rolled, then the probability of obtaining at least two sixes is equal to $\qquad$ (round off to 2 decimal places) |
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| Q. 45 | Let $X$ be a random variable with the moment generating function $M(t)=\frac{\left(1+3 e^{t}\right)^{2}}{16},-\infty<t<\infty$ <br> Let $\alpha=E(X)-\operatorname{Var}(X)$. Then the value of $8 \alpha$ is equal to $\qquad$ (answer in integer) |
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| Q. 46 | For $n \in \mathbb{N}$, let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample from the Cauchy distribution having probability density function $f(x)=\frac{1}{\pi\left(1+x^{2}\right)},-\infty<x<\infty .$ <br> Let $g: \mathbb{R} \rightarrow \mathbb{R}$ be defined by $g(x)=\left\{\begin{array}{cc} x, & \text { if }-1000 \leq x \leq 1000 \\ 0, & \text { otherwise } \end{array}\right.$ <br> Let $\alpha=\lim _{n \rightarrow \infty} P\left(\frac{1}{n^{\frac{3}{4}}} \sum_{i=1}^{n} g\left(X_{i}\right)>\frac{1}{2}\right) .$ <br> Then $100 \alpha$ is equal to $\qquad$ (answer in integer) |


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| Q. 47 | For $n \in \mathbb{N}$, let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample from the $F_{20,40}$ distribution. Then, as $n \rightarrow \infty, \frac{1}{n} \sum_{i=1}^{n} \frac{1}{X_{i}}$ converges in probability to $\qquad$ (round off to 2 decimal places) |
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| Q. 48 | Let $X_{1}, X_{2}, \ldots, X_{10}$ be a random sample from the $\operatorname{Exp}(1)$ distribution. Define $W=\max \left\{e^{-X_{1}}, e^{-X_{2}}, \ldots, e^{-X_{10}}\right\}$. Then the value of $22 E(W)$ is equal to $\qquad$ (answer in integer) |
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| Q. 49 | Let $X_{1}, X_{2}, X_{3}$ be i.i.d. random variables from a continuous distribution having probability density function $f(x)=\left\{\begin{array}{cl} \frac{1}{2 x^{3}}, & \text { if } x>\frac{1}{2} \\ 0, & \text { if } x \leq \frac{1}{2} \end{array}\right.$ <br> Let $X_{(1)}=\min \left\{X_{1}, X_{2}, X_{3}\right\}$. Then the value of $10 E\left(X_{(1)}\right)$ is equal to $\qquad$ (answer in integer) |


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| Q.50 | Suppose that the lifetimes (in months) of bulbs manufactured by a company have <br> an $\operatorname{Exp}(\lambda)$ distribution, where $\lambda>0$. A random sample of size 10 taken from the <br> bulbs manufactured by the company yields the sample mean lifetime $\bar{x}=3.52$ <br> months. Then the uniformly minimum variance unbiased estimate of $\frac{1}{\lambda}$ based on <br> this sample is equal to <br> months (round off to 2 decimal places) |
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Section C: Q. 51 - Q. 60 Carry TWO marks each.

| Q. 51 | The value of $\lim _{n \rightarrow \infty} n\left(\sin \frac{1}{2 n}-\frac{1}{2} e^{-\frac{1}{n}}+\frac{1}{2}\right)$ is equal to _____ (answer in <br> integer) |
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| Q.52 | The value of the integral <br> is equal to $\quad \int_{0}^{2} \int_{x}^{\sqrt{8-x^{2}}} \frac{3 \sqrt{x^{2}+y^{2}}}{\sqrt{8} \pi} \mathrm{~d} y \mathrm{~d} x$ |
| (answer in integer) |  |


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| Q.53 | For some $a \leq 0$ and $b \in \mathbb{R}$, let |
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| Q. 55 | Let $X$ be a discrete random variable with $P(X \in\{-5,-3,0,3,5\})=1$. <br> Suppose that $\begin{gathered} P(X=-3)=P(X=-5), \\ P(X=3)=P(X=5) \text { and } \\ P(X>0)=P(X=0)=P(X<0) . \end{gathered}$ <br> Then the value of $12 P(X=3)$ is equal to $\qquad$ (answer in integer) |
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| Q. 56 | Consider a coin for which the probability of obtaining head in a single toss is $\frac{1}{3}$. Sunita tosses the coin once. If head appears, she receives a random amount of $X$ rupees, where $X$ has the $\operatorname{Exp}\left(\frac{1}{9}\right)$ distribution. If tail appears, she loses a random amount of $Y$ rupees, where $Y$ has the $\operatorname{Exp}\left(\frac{1}{3}\right)$ distribution. Her expected gain (in rupees) is equal to $\qquad$ (answer in integer) |
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| Q. 57 | Let $\Theta$ be a random variable having $U(0,2 \pi)$ distribution. Let $X=\cos \Theta$ and <br> $Y=\sin \Theta$. Let $\rho$ be the correlation coefficient between $X$ and $Y$. Then $100 \rho$ is <br> equal to (answer in integer) |
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| Q.58 | Let $X_{1}, X_{2}, \ldots, X_{10}$ be a random sample from a $U(-\theta, \theta)$ distribution, where <br> $\theta \in(0, \infty)$. Let $X_{(10)}=\max \left\{X_{1}, X_{2}, \ldots, X_{10}\right\}$ and $X_{(1)}=$ min $\left\{X_{1}, X_{2}, \ldots, X_{10}\right\}$. If <br> the observed values of $X_{(10)}$ and $X_{(1)}$ are 8 and -10, respectively, then the <br> maximum likelihood estimate of $\theta$ is equal to |
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| Q. 59 | Suppose that the weights (in kgs ) of six months old babies, monitored at a healthcare facility, have $N\left(\mu, \sigma^{2}\right)$ distribution, where $\mu \in \mathbb{R}$ and $\sigma>0$ are unknown parameters. Let $X_{1}, X_{2}, \ldots, X_{9}$ be a random sample of the weights of such babies. Let $\bar{X}=\frac{1}{9} \sum_{i=1}^{9} X_{i} \quad, \quad S=\sqrt{\frac{1}{8} \sum_{i=1}^{9}\left(X_{i}-\bar{X}\right)^{2}} \quad$ and let a $95 \%$ confidence interval for $\mu$ based on $t$-distribution be of the form $(\bar{X}-h(S), \bar{X}+h(S)),$ <br> for an appropriate function $h$ of random variable $S$. If the observed values of $\bar{X}$ and $S^{2}$ are 9 and 9.5 , respectively, then the width of the confidence interval is equal to $\qquad$ (round off to 2 decimal places) <br> $\left(\right.$ You may use $\left.t_{9,0.025}=2.262, t_{8,0.025}=2.306, t_{9,0.05}=1.833, t_{8,0.05}=1.86\right)$ |
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| Q. 60 | Let $X_{1}, X_{2}, X_{3}$ be a random sample from a Poisson distribution with mean $\lambda, \lambda>0$. For testing $H_{0}: \lambda=\frac{1}{8}$ against $H_{1}: \lambda=1$, a test rejects $H_{0}$ if and only if $X_{1}+X_{2}+X_{3}>1$. The power of this test is equal to $\qquad$ (round off to 2 decimal places) |

